

RESEARCH

Open Access

Translation and validation of the EORTC brain cancer module (EORTC QLQ-BN20) for use in Iran

Alireza Khoshnevisan¹, Mir Saeed Yekaninejad^{2,5}, Shahab Kamali Ardakani¹, Amir H Pakpour^{3,4*}, Azam Mardani⁵ and Neil K Aaronson⁶

Abstract

Background: The aim of this study was to translate the EORTC quality of life questionnaire for brain cancer, the QLQ-BN20, into Persian, and to evaluate its psychometric properties when used among brain cancer patients in Iran.

Methods: A standard backward and forward translation procedure was used to generate the Persian language version of the QLQ-BN20. The QLQ-BN20 was administered together with the QLQ-C30 to 194 patients diagnosed with primary brain cancer. Multitrait scaling and confirmatory factor analysis (CFA) were used to evaluate the hypothesized scale structure of the questionnaire. Internal consistency reliability was estimated with Cronbach's alpha. The ability of the QLQ-BN20 to distinguish between patient subgroups formed on the basis of performance status and cognitive status was evaluated, as was the responsiveness of the questionnaire to changes in performance status over time.

Results: Multitrait scaling and CFA results confirmed the hypothesized scale structure. The measurement model was consistent across men and women. Internal consistency reliability of the multi-item scales ranged from 0.74 to 0.89. The QLQ-BN20 distinguished clearly between patients with relatively good versus poor performance and cognitive status, and changes in scores over time reflected changes observed in performance status ratings.

Conclusions: These results support the validity and reliability of the QLQ-BN20 for use among Iranian patients diagnosed with primary brain cancer. Future studies should examine the psychometrics of the questionnaire when used in patients with brain metastasis.

Keywords: Quality of life, Brain cancer, QLQ-BN20, QLQ-C30, Psychometric evaluation

Background

It is estimated that 22,020 patients are diagnosed annually with brain cancer in the United States [1]. In Iran, brain cancer is the 11th most common cancer, with an annual incidence rate of approximately 550 cases for males and 800 for females [2]. Although brain cancer is relatively rare, it is a disease with serious symptoms and a poor prognosis. [1,3]. The 5-year relative survival rate is 36% [1].

There is no definite cure for patients with brain cancer. Therefore, the primary aim of treatment is to

prolong the patient's life and to palliate symptoms [4]. The treatment for the most common brain tumors, gliomas, includes neurosurgery, followed by adjuvant radiotherapy, chemotherapy or combined chemoradiotherapy. All of these treatments may cause significant complications or have toxic side-effects [5,6]. Therefore, in order to select the best treatment for a specific patient it is necessary to take into consideration the stage of the disease, any risk associated with the given treatment and the patient's general health. An assessment of patient-reported outcomes, and more specifically health-related quality of life (HRQOL), is very important in this respect.

Two well-known, generic questionnaires for assessing the HRQOL of patients with cancer include The European Organization for Research and Treatment of Cancer Core Questionnaire (the EORTC QLQ-C30) and the

* Correspondence: Pakpour_Amir@yahoo.com

³Qazvin Research Center for Social Determinants of Health, Qazvin University of Medical Sciences, Qazvin, Iran

⁴Department of Public Health, Qazvin University of Medical Sciences, Qazvin, Iran

Full list of author information is available at the end of the article

Functional Assessment of Cancer Therapy-General (FACT-G) [7,8]. These generic questionnaires are applicable to all types of patients with cancer. Both the EORTC and the FACT measurement systems supplement their core questionnaire with condition-specific questionnaires designed to capture functional limitations and symptoms experienced by specific populations of patients. Both the EORTC and the FACT systems have developed questionnaires specific to brain cancer. The EORTC brain cancer-specific questionnaire, the QLQ-BN20, has been validated and has been translated into a number of primarily European languages [9,10]. The aim of the present study was to translate the QLQ-BN20 into the dominant language in Iran, Persian (spoken by approximately two-thirds of the population), and to generate psychometric data regarding the questionnaires reliability and validity when used among Iranian brain cancer patients.

Methods

Translation of the QLQ-BN20

The QLQ-BN20 questionnaire contains 20 items organized into four scales; future uncertainty (3 items), visual disorder (3 items); motor dysfunction (3 items); and communication deficit (3 items), and seven single items (headaches, seizures, drowsiness, hair loss, itchy skin, weakness of legs, and bladder control) [9,10]. All items are rated on a four-point Likert-type scale ('not at all', 'a little', 'quite a bit' and 'very much'), and are linearly transformed to a 0-100 scale, with higher scores indicating more severe symptoms [11].

We obtained permission from the EORTC Quality of Life Department to generate the Iranian version of the QLQ-BN20. The aim of the translation procedure was to provide an Iranian version of the questionnaire that was conceptually equivalent to the original English language version.

We used the standard EORTC translation procedures [12]. Briefly, two trained bilingual Iranians independently translated the QLQ-BN20 into Persian. A specialized translator (project manager; AHP) compared these two versions, evaluated the level of difficulty in terms of translation as well as the equivalence of each item and response scale. The translators compared their translations, reconciled their discrepancies and arrived at a unified Persian version of the questionnaire. Subsequently, the provisional Persian version was independently translated back into U.S. English by two additional translators who were native English speakers. The Project Manager compared the English translations with the original questionnaire, and any discrepancies between the back-translations were resolved with the translators. This process went through two iterations, and the documentation was reviewed by the EORTC Quality of Life

Group. The resulting Persian version of the QLQ-BN20 was administered to 10 brain tumor patients (6 women and 4 men; age range = 28–69 years). All patients were able to complete the questionnaire without assistance, and reported that the questions were clear and easy to complete.

Validation of the Persian version of the QLQ-BN20

Study participants were patients with primary brain cancer who had been referred to Shariati Hospital and the Cancer Institute of Iran in Tehran. These centers are national referral centers for cancer treatment in Iran. Inclusion criteria were; a histologically verified diagnosis of brain cancer, older than 18 years, fluency in Persian, and a life expectancy of at least 4 weeks. All patients provided written informed consent. Patients with a previous cancer diagnosis, those who were too frail and those who were unable to understand the questionnaire or unwilling to provide written informed consent, were excluded. The study protocol was approved by the Ethical Committee of Tehran University of Medical Sciences. All participants gave their written informed consent.

Data collection

Patients were recruited into the study after surgery and before chemotherapy (CT), radiotherapy (RT) or chemoradiotherapy (CRT). Patients completed the validated Persian version of the QLQ-C30 [13] and the Persian version of the QLQ-BN20 before the start of the CT, RT or CRT, and again approximately four weeks later. We obtained sociodemographic and clinical information via questionnaire and medical record review. This included age, sex, education, tumor type, and surgery. Additionally, the Mini-Mental State Examination (MMSE) and The Karnofsky Performance Status Scale (KPS) were administered to all patients.

The questionnaire was completed by interview if the patients were illiterate. The interviewer read the questions to the patients without embellishment or explanation, and presented the possible responses and elicited patients' choices without coaching.

Statistical analysis

Ceiling and floor effects were calculated for all QLQ-BN20 scales. A ceiling or floor effect lower than 15% indicates acceptable measurement standards [14].

We evaluated the scale structure of the QLQ-BN20 in several ways. First, we used multitrait scaling to assess item convergent and discriminant validity. We calculated Spearman correlation coefficients between each item and its own scale, corrected for overlap. Item-scale correlations of 0.40 or greater were taken as evidence of item convergent validity [15]. We also calculated the

correlations between each item and the other scales. We expected that items would correlate more highly (two times the standard error, $1/\sqrt{N}$) with their own scale than with other scales.

Second, we used confirmatory factor analysis (CFA) to test the hypothesized model fit. We employed three categories of indices: absolute fit indices, incremental fit indices and parsimony fit indices. An absolute fit index assesses how well an *a priori* model reproduces the sample data. We used chi-square, root mean square error of approximation (RMSEA) and standardized root mean square residual (SRMR) as absolute fit indices [16,17]. Chi-square assesses the magnitude of discrepancy between the sample and fitted covariance matrices [18]. However, the chi-square statistic is sensitive to sample size and therefore with large samples, is not a practical test of model fit. RMSEA incorporates a penalty function for poor model parsimony [16]. RMSEA, values in the range of 0.05–0.08 were taken to indicate acceptable fit, values in the range of 0.08–0.10 to indicate marginal fit, and values larger than 0.10 to indicate poor fit [16]. SRMR is the square root of the difference between the residuals of the sample covariance matrix and the hypothesized covariance model. Values of the SRMR less than 0.10 are generally considered favorable.

We used the non-normed fit index (NNFI, also known as the Tucker-Lewis index) and the comparative fit index (CFI) to estimate incremental fit. The suggested cut off for NNFI and CFI is 0.90 or greater.

We computed the parsimonious normed fit index (PNFI) to assess the parsimony the model [17,19]. For PNFI we did not employ any absolute standard of model fit, but rather simply noted that higher PNFI values reflect more parsimonious fit [19]. We used weighted least squares (WLS) as the method of estimation.

In addition, we examined the model invariance across gender. To compare the factor loadings across gender, we applied multi-group measurement invariance analysis. There are different terminologies in the literature for tests of invariance. We used two terms of factorial invariance (i.e., configural invariance and metric invariance). In configural invariance, the pattern of fixed- and free-factor loadings is constant across groups, while the magnitudes of these loadings are not constrained to be equal. For metric invariance, the magnitudes of factor loadings for particular items are invariant across groups [20]. As suggested by Cheung, differences in practical fit indices such as CFI and NNFI not larger than 0.01 between NNFI or CFI values were considered as evidence of model invariance [20]. We hypothesized that the Persian version of QLQ-BN20 would perform similarly across gender.

We estimated internal consistency reliability of the QLQ-BN20 scales with Cronbach's coefficient alpha. An

alpha coefficient of 0.70 or higher was considered acceptable for purposes of group comparisons [21].

“Known groups” or clinical validity was evaluated by comparing the QLQ-BN20 scores between patients grouped according to KPS and MMSE scores. We hypothesized that patients with higher KPS (>80) would score better on the QLQ-BN20 than those with lower KPS (≤ 80) [9,10]. We also anticipated that patients with

Table 1 Sociodemographic and clinical characteristics of the study sample (n = 194)*

	Total
Age (years)Mean (SD)	42.5 (16.1)
Sex	
Male	103 (53.0%)
Female	91 (47.0%)
Educational status	
Illiterate	35 (18.0%)
Primary school	33 (17.0%)
Middle school	36 (18.6%)
Secondary school	48 (24.8%)
College	42 (21.6%)
Marital status	
Single	43 (22.2%)
Married	128 (65.9%)
Widowed/divorced	23 (11.9%)
KPS	
≤ 80	102 (52.5%)
>80	92 (47.5%)
MMSE	
Mean (SD)	23.89(5.3)
Median	26.00
Range	8-30
Type of surgery	
Biopsy only	40 (20.6%)
Partial resection	96 (49.5%)
Total resection	58 (29.9%)
Type of adjuvant therapy	
CT	75(38.7%)
RT	62 (32%)
CT+ RT	57 (29.3%)
Tumor type	
Astrocytoma	72(37.1%)
Atypical Meningioma*	46(23.7%)
Oligodendroglioma	53(27.3%)
Others	23(11.9%)

KPS = Karnofsky Performance Status; MMSE = Mini-Mental State Examination.

*Patients with grade 2 and 3 tumors including, chordoid, atypical, clear cell, rhabdoid, anaplastic (malignant) and papillary subtypes. Meningeal sarcomas (angioblastic meningiomas) and hemangiopericytomas were also included in the Atypical Meningioma.

higher MMSE (≥ 27) would score better on the QLQ-BN20 than those with lower MMSE (< 27) [10]. Analysis of covariance (ANCOVA) was used to test for group differences. To control for probability of type I errors due to multiple comparisons, we used an adjustment procedure developed by Benjamini and Hochberg. This procedure controls the false discovery rate. The false discovery rate level was set at 5% [22,23]. Effect sizes (Cohen's d statistic) were calculated to estimate the magnitude of observed, statistically significant group differences [24].

Finally, to evaluate the responsiveness of the QLQ-BN20 to change in health status over time, we classified patients as having worse, stable or improved KPS scores from baseline to follow-up. We evaluated changes in QLQ-BN20 scores as a function of change in KPS with analysis of covariance (ANCOVA), adjusting for baseline values.

We used SPSS 16.0 for Windows and LISREL 8.8 for data analyses. A p value < 0.05 was considered as statistically significant.

Results

In total, 194 patients were recruited into the study. The patients' baseline characteristics are provided in Table 1. The mean age of the patients was 42 ± 5 years, with a

range between 18 and 80 years. Forty-seven percent of the patients was female. The majority of patients (66%) was married, and 46% had completed secondary school or college.

Astrocytomas (37.1%) constituted the most frequent tumor type. Slightly more than one-third of patients were receiving CT, one-third RT, and slightly less than one-third CT + RT. Only seven patients (3.65%) were lost to follow-up assessment due to death.

As shown in Table 2, all subscales of the QLQ-BN20 were found to be normally distributed (Kolmogorov-Smirnov, $p > 0.05$). There was no ceiling effect for any of the BN20 subscales. However, some floor effects were observed for seizures, headaches, itchy skin and bladder control.

Results of the multitrait scaling analysis are presented in Table 3. All items correlated 0.60 or higher with their own scale, corrected for overlap. There were no scaling errors, with all items correlating two standard errors or higher with their own scales than with other scales.

Results of the CFA analysis indicated that all of the goodness of fit indices supported the four-factor model for the QLQ-BN20: $\chi^2 = 92.51$, degree of freedom = 59, $p < 0.001$, CFI = 1.0, RMSEA = 0.068, SRMR = 0.066, NNFI = 1.0, and PNFI = 0.73. Moreover, all factor

Table 2 Descriptive statistics for the QLQ-BN20

		Number of forms	Mean (SD)	Floor N (%)	Ceiling (%)	Cronbach's alpha	Normality (Kolmogoroff-Smirnoff-Test)
BFU (future uncertainty)	Baseline	194	39.9 (24.9)	8 (4.1%)	3 (1.5%)	0.80	0.14
	Follow-up	187	36.8 (24.5)	10 (5.3%)	3 (1.6%)	0.74	0.24
BVD (visual disorder)	Baseline	194	27.9 (26.9)	17 (8.7%)	5 (2.6%)	0.74	0.16
	Follow-up	187	25.8 (25.9)	14 (7.4%)	4 (2.1%)	0.80	0.15
BMD (motor dysfunction)	Baseline	194	29.0 (28.9)	28 (14.4%)	6 (3.1%)	0.80	0.46
	Follow-up	186	28.6 (29.6)	23 (12.3%)	3 (1.6%)	0.82	0.18
BCD (communication deficit)	Baseline	194	23.7 (27.2)	18 (9.2%)	3 (1.5%)	0.89	0.21
	Follow-up	187	20.9 (25.0)	14 (7.5%)	1 (0.5%)	0.89	0.22
BHA (headaches)	Baseline	192	46.9 (33.9)	25 (13.0%)	24 (12.5%)	-	0.19
	Follow-up	187	43.6 (35.1)	21 (11.2%)	18 (9.6%)	-	0.15
BSE (seizures)	Baseline	194	10.0 (21.6)	99 (51.5%)	7 (3.6%)	-	0.13
	Follow-up	186	7.7 (20.0)	95 (51.0%)	1 (0.5%)	-	0.11
BDR (drowsiness)	Baseline	192	32.2 (31.7)	28 (14.6%)	11 (5.7%)	-	0.6
	Follow-up	185	31.2 (32.0)	19 (10.3%)	8 (4.3%)	-	0.5
BHL (hair loss)	Baseline	194	23.8 (32.5)	41 (21.1%)	9 (4.6%)	-	0.28
	Follow-up	187	25.9 (44.4)	38 (20.3%)	10 (5.3%)	-	0.06
BIS (itchy skin)	Baseline	193	14.0 (25.8)	54 (28.0%)	5 (2.6%)	-	0.20
	Follow-up	186	17.8 (27.5)	46 (24.7%)	6 (3.2%)	-	0.32
BWL (weakness of legs)	Baseline	190	28.4 (33.9)	23 (12.1%)	14 (7.4%)	-	0.36
	Follow-up	187	28.3 (34.1)	21 (11.2%)	12 (6.4%)	-	0.29
BBC (bladder control)	Baseline	194	15.1 (25.8)	67 (34.5%)	4 (2.0%)	-	0.20
	Follow-up	186	15.2 (30.3)	58 (31.2%)	6 (3.2%)	-	0.30

Table 3 Results of the multitrait scaling analysis*

	BFU	BVD	BMD	BCD
BFU (future uncertainty)				
uncertain about the future	0.75	0.15	0.04	0.19
setbacks in condition	0.62	0.06	0.10	0.28
disruption of family life	0.78	0.23	0.10	0.17
future worsen	0.67	0.21	0.18	0.04
BVD (visual disorder)				
double vision	0.14	0.67	0.11	0.18
vision blurred	0.08	0.86	0.02	0.07
difficulty reading	0.16	0.83	0.10	0.05
BMD (motor dysfunction)				
weakness on one side of body	0.10	0.25	0.79	0.15
trouble with coordination	0.01	0.12	0.80	0.19
feel unsteady on your feet	0.18	0.12	0.86	0.18
BCD (communication deficit)				
Trouble finding the right words to express yourself	0.14	0.21	0.11	0.88
Difficulty speaking	0.13	0.13	0.12	0.88
Trouble communicating thoughts	0.09	0.23	0.13	0.87

* Correlations are based on analysis of the baseline data.

loadings were significant, ranging from 0.24 to 0.89 (Figure 1). Additional analyses were based on a four-factor model that constrained the factor loadings to be equal across gender groups. The measurement model revealed an acceptable model fit for the QLQ-BN20: $\chi^2 = 190.5$, degree of freedom = 133, $p < 0.001$, CFI = 0.923, SRMR = 0.089, NNFI = 0.911, PNFI = 0.704, RMSEA = 0.080. Moreover, metric invariance showed good fit indices: $\chi^2 = 205.15$, degree of freedom = 133, $p < 0.001$, CFI = 0.927, SRMR = 0.091, NNFI = 0.915, PNFI = 0.753, RMSEA = 0.079. These results support the stability of the hypothesized factor structure across sexes.

The internal consistency reliability of all scales of the QLQ-BN20 was acceptable for group comparisons (Table 2). Cronbach's alpha coefficient for all scales was above the 0.70 threshold. Communication deficit had the highest internal consistency (0.89 at both baseline and follow-up) while the lowest was for future uncertainty (0.80, 0.74).

Results from the known group comparisons between subgroups of patients formed on the basis of performance status and MMSE are summarized in Table 4. As hypothesized, patients with better KPS and MMSE

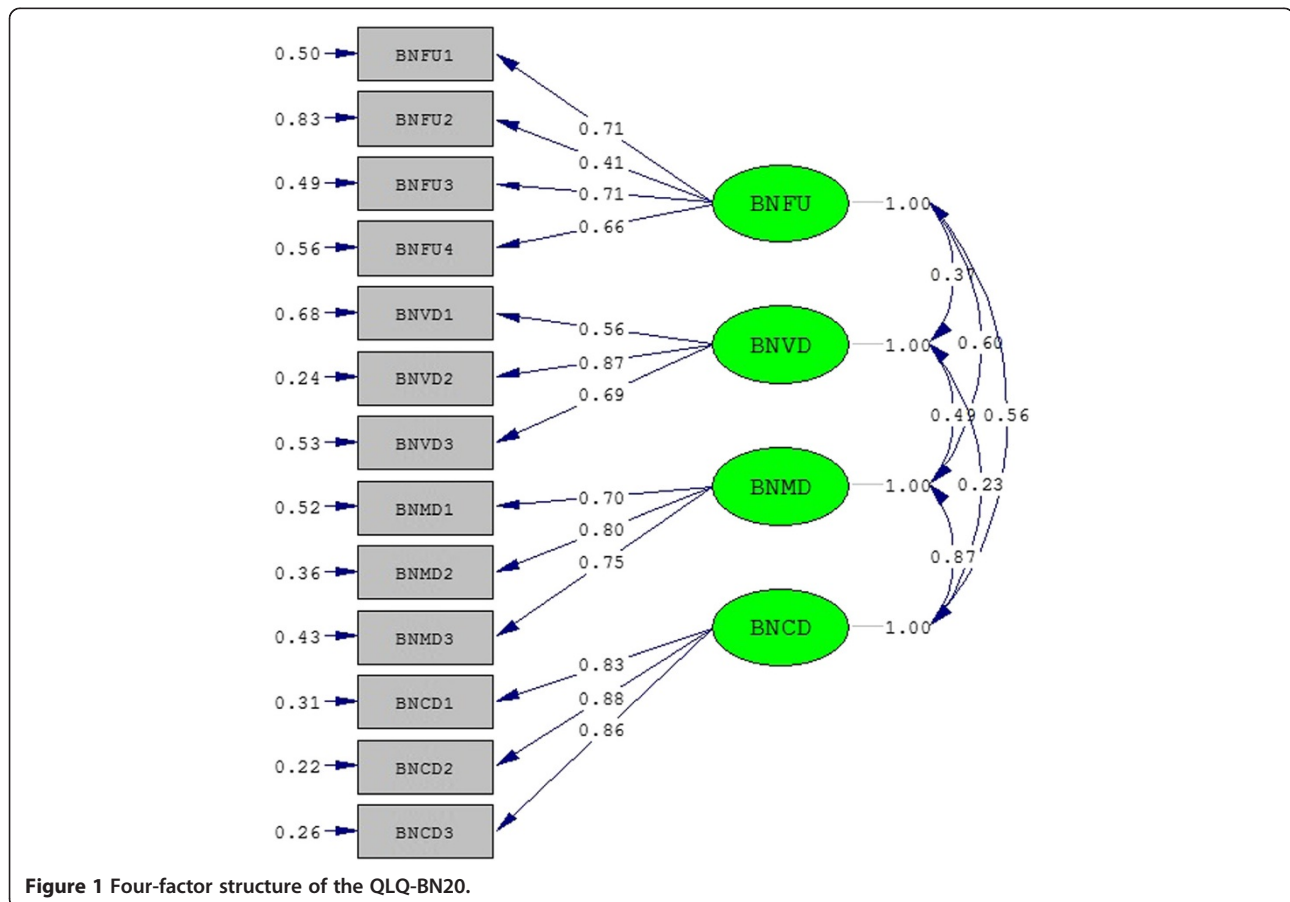


Figure 1 Four-factor structure of the QLQ-BN20.

Table 4 Known groups validity testing: Comparison of QLQ-BN20 Scores as a function of KPS and MMSE

	Baseline Follow-up	KPS		Effect size	MMSE**		Effect size
		>80	≤80		<27	≥27	
		N = 53, N = 41	N = 141, N = 153		N = 91	N = 103	
BFU (future uncertainty)	Baseline*‡	42.7 (22.2)	33.5 (24.7)	0.38	33.9 (24.4)	27.2 (23.4)	0.28
	Follow-up *	37.0 (15.1)	25.6 (24.5)	0.50	-	-	-
BVD (visual disorder)	Baseline *‡	38.3 (30.1)	27.0 (27.3)	0.40	28.1 (29.4)	21.8 (22.8)	0.24
	Follow-up *	33.3 (27.8)	23.7 (26.1)	0.36	-	-	-
BMD (motor dysfunction)	Baseline*‡	68.2 (31.8)	24.6 (26.3)	1.56	33.7 (30.2)	16.5 (22.5)	0.65
	Follow-up *	63.9 (29.1)	22.4 (27.0)	1.5	-	-	-
BCD (communication deficit)	Baseline *‡	44.9 (31.9)	20.3 (26.5)	0.88	26.1 (28.3)	14.1 (21.8)	0.48
	Follow-up*	45.0 (33.4)	17.0 (22.6)	1.11	-	-	-
BHA (headaches)	Baseline *‡	50.0 (36.0)	44.3 (32.2)	0.17	47.8 (28.6)	34.8 (34.4)	0.41
	Follow-up *	63.3 (36.7)	42.5 (35.1)	0.59	-	-	-
BSE (seizures)	Baseline *‡	15.1 (22.9)	11.0 (23.3)	0.18	14.3 (25.5)	5.7 (15.6)	0.41
	Follow-up *	14.8 (24.2)	8.5 (21.2)	0.29	-	-	-
BDR (drowsiness)	Baseline *‡	57.6 (30.1)	28.3 (31.2)	0.95	34.7 (33.3)	22.3 (25.0)	0.42
	Follow-up *	66.7 (27.2)	26.6 (30.5)	1.34	-	-	-
BHL (hair loss)	Baseline*‡	33.3 (35.1)	22.1 (32.2)	0.34	23.4 (32.8)	16.8 (33.5)	0.20
	Follow-up*	63.3 (36.7)	42.5 (35.1)	0.59	-	-	-
BIS (itchy skin)	Baseline*‡	33.3 (35.1)	13.2 (24.7)	0.72	19.2 (30.4)	9.2 (17.6)	0.41
	Follow-up*	40.0 (37.8)	16.5 (26.5)	0.80	-	-	-
BBC (bladder control)	Baseline*‡	45.4 (37.3)	12.8 (24.2)	1.15	19.0 (27.3)	9.2 (23.4)	0.39
	Follow-up*	50.0 (42.3)	12.9 (28.2)	0.97	-	-	-

* Statistically significant for KPS.
 ‡ Statistically significant for MMSE.
 ** MMSE was only available at baseline.

scores had significantly better scores on all of the QLQ-BN20 multi-item scales than patients with lower KPS and MMSE scores. Effect sizes (Cohen's d statistic) ranged from 0.18 to 1.56 for the KPS, and from 0.20 to 0.65 for the MMSE.

Finally, changes in QLQ-BN20 scores over time as a function of changes in KPS scores (worse, stable or improved performance status) are reported in Table 5. The results supported the ability of all of the QLQ-BN20 scales to detect such shifts in KPS, with the exception of seizures, drowsiness, visual disorder and hair loss which did not vary as a function of change in KPS.

Discussion

In this paper we have reported on the results of the translation and validation of the Persian version of the QLQ-BN20 for use with primary brain cancer patients in Iran. This included examination of the hypothesized scale structure of the questionnaire, internal consistency reliability estimation, known groups validity testing, and evaluation of the responsiveness of the QLQ-BN20 to changes over time in health status.

Overall, the results lend strong support to the psychometric robustness of the questionnaire. The large majority of patients completed all items and there were very minor missing items. Ceiling effects were absent while

Table 5 Changes of QLQ-BN20 over time by performance status

	KSP			P value (ANCOVA)
	Worsening	Stable	Improved	
	N = 59 Mean (SD)	N = 49 Mean (SD)	N = 79 Mean (SD)	
Headaches *	12.4 (29.2)	1.2 (35.4)	-10.0 (32.6)	0.04
Motor dysfunction*	6.1 (18.8)	0.1 (18.2)	-4.3 (15.4)	0.01
Drowsiness	5.5 (23.2)	2.4 (27.3)	1.7 (24.3)	0.38
Weakness of legs ^{††}	15.2 (28.4)	-1.3 (20.2)	-12.1 (19.3)	<0.01
Visual disorder	5.4 (16.1)	2.7 (25.4)	1.6 (22.3)	0.64
Bladder control ^{††}	9.6 (18.8)	-0.7 (28.9)	-6.6 (26.3)	0.01
Seizures	5.3 (19.8)	3.2 (21.8)	2.7 (17.6)	0.72
Hair loss	12.4 (29.2)	-4.5 (50.4)	-15.2 (45.4)	0.08
Itchy skin*	6.5 (20.8)	0.4 (26.4)	-3.8 (24.8)	0.02

* P < 0.05 Improved Vs. Worsening Group.
 † P < 0.05 Stable Vs. Worsening Group.

some floor effects were observed for items assessing seizures, hair loss, itchy skin and bladder control. A possible reason for this latter finding could be that most of the patients were newly-diagnosed and had only recently initiated adjuvant treatment. With longer follow-up, more variance in scores for these symptoms could be expected. Our results are in accordance with previous studies [9,10].

The results of both the multitrait scaling and CFA provided strong support for the hypothesized scale structure of the QLQ-BN20. To our knowledge, ours is the first study to employ CFA in the evaluation of the QLQ-BN20 measurement model. Some studies have reported that females with brain tumors tend to report worse QOL than their male counterparts [25-27]. It is important to determine if these gender differences also affect the scale structure of the questionnaire. [28]. Toward this end, we assessed the factorial invariance of the QLQ-BN20. The results of the CFA demonstrated that the hypothesized four-factor model fit the data of both the male and female samples well, supporting invariance. To our knowledge, ours is the first study to assess factor structure of the QLQ-BN20 for patients with primary brain tumours using confirmatory factor analysis. We would recommend that this be undertaken in other cultural and language settings as well.

The internal consistency reliability of the QLQ-BN20 scales exceeded the recommended 0.70 level, with most coefficients in the 0.80 to 0.90 range. These reliability estimates are even higher than those reported in previous studies with English-speaking and multi-national samples [9,10].

The QLQ-BN20 was able to distinguish clearly between subgroups of patients formed on the basis of their performance status and cognitive status, and was responsive to change over time in performance status. Again, these results are in line with those of previous studies [9,10].

Our study had several limitations that should be noted. First, we recruited a convenience sample of patients. Nevertheless, we believe that the sample sufficiently represents the population of interest. Second, although the overall sample size was sufficient for the analyses that we conducted, there were insufficient numbers available to carry out subgroup analyses of more homogenous groups of brain cancer patients. This is something that could be done in future studies. Finally, the study was of relatively short duration, and thus we had a relatively short period to evaluate the responsiveness of the QLQ-BN20 over time. Use of the questionnaire in large, prospective studies should provide the opportunity to examine the responsiveness of the questionnaire over long periods of time.

In conclusion, the results of our study support the feasibility, measurement model, reliability and validity

of the QLQ-BN20 in assessing the HRQOL of Iranian patients with primary brain cancer. The availability of this questionnaire in Persian will facilitate the assessment of the health-related quality of life of brain cancer patients in Iran, particularly in the context of clinical research, but also potentially in clinical practice setting. Future studies are needed to examine the psychometric properties of the Iranian version of the QLQ-BN20 among patients with brain metastases.

Competing interests

The authors declare that they have no competing interests.

Author details

¹Department of neurosurgery, Shariati hospital, Tehran University of medical sciences, Tehran, Iran. ²Department of Epidemiology and Biostatistics, School of Public Health, Tehran University of Medical Sciences, Tehran, Iran. ³Qazvin Research Center for Social Determinants of Health, Qazvin University of Medical Sciences, Qazvin, Iran. ⁴Department of Public Health, Qazvin University of Medical Sciences, Qazvin, Iran. ⁵Brain and Spinal Injury Research Center (BASIR), Imam Khomeini Hospital Complex, Tehran University of medical sciences, Tehran, Iran. ⁶Division of Psychosocial Research & Epidemiology, The Netherlands Cancer Institute, Plesmanlaan 1211066 CX, Amsterdam, the Netherlands.

Authors' contributions

AKh and AHP were responsible for designing the study, analyzing and interpreting the data, and drafting the manuscript. ShKA and AM were responsible for data collection. NKA interpreted the data, and revised the manuscript. The analyses were performed by MY. All authors have read and approved the final manuscript.

Received: 21 February 2012 Accepted: 20 May 2012

Published: 20 May 2012

References

1. Jemal A, Siegel R, Xu J, Ward E: **Cancer statistics, 2010**. *CA: a cancer journal for clinicians* 2010, **60**(5):277-300.
2. Iran center for disease control and prevention, cancer office: *Iranian annual of national cancer registration report 2006-2007*. Tehran: Ministry of health; 2008.
3. US News & World Report: Brain Tumor Channel: *Cleveland Clinic [cited 2012 May]*: Available from: URL: <http://health.usnews.com/health-conditions/cancer/brain-tumor>.
4. Kirkbride P, Tannock IF: **Trials in palliative treatment – have the goal posts been moved?** *Lancet Oncology* 2008, **9**:186-187.
5. Stupp R, Mason WP, van den Bent MJ, et al: **Radiotherapy plus concomitant and adjuvant temozolomide for glioblastoma**. *N Engl J Med* 2005, **352**:987-996.
6. Brandsma D, van den Bent MJ: **Molecular targeted therapies and chemotherapy in malignant gliomas**. *Curr Opin Oncol* 2007, **19**(6):598-605.
7. Aaronson NK, Ahmedzai S, Bergman B, et al: **The European Organization for Research and Treatment of Cancer QLQ-C30: a quality-of-life instrument for use in international clinical trials in oncology**. *J Natl Cancer Institute* 1993, **85**(5):365-376.
8. Cella DF, Tulsky DS, Gray G, et al: **The Functional Assessment of Cancer Therapy Scale: development and validation of the general measure**. *J oncology: official j Am Soc Clin Oncology* 1993, **11**(3):570-579.
9. Osoba D, Aaronson NK, Muller M, et al: **The development and psychometric validation of a brain cancer quality-of-life questionnaire for use in combination with general cancer specific questionnaires**. *Quality of Life Res* 1996, **5**(1):139-150.
10. Taphoorn MJ, Claassens L, Aaronson NK, Coens C, Mauer M, Osoba D, Stupp R, Mirimanoff RO, van den Bent MJ, Bottomley A: **EORTC Quality of Life Group, and Brain Cancer, NCIC and Radiotherapy Groups. An international validation study of the EORTC brain cancer module (EORTC QLQ-BN20) for assessing health-related quality of life and symptoms in brain cancer patients**. *Eur J Cancer* 2010, **46**(6):1033-1040.

11. Fayers PM, Aaronson NK, Bjordal K, Groenvold M, Curran D, Bottomley A: *on behalf of the EORTC Quality of Life Group, The EORTC QLQ-C30 Scoring Manual*. 3rd edition. Brussels: EORTC; 2001.
12. Cull A, Sprangers MAG, Aaronson NK: *On behalf of the EORTC Quality of Life Study Group. Internal report of the EORTC Study Group on Quality of Life: European Organization for Research and Treatment of Cancer quality of life study group translation procedure*; 1994.
13. Montazeri A, Harirchi I, Vahdani M, Khaleghi F, Jarvandi S, Ebrahimi M, Haji-Mahmoodi M: **The European Organization for Research and Treatment of Cancer Quality of Life Questionnaire (EORTC QLQ-C30): translation and validation study of the Iranian version.** *Supportive Care in Cancer* 1999, **7**(6):400–406.
14. Terwee CB, Bot SD, de Boer MR, et al: **Quality criteria were proposed for measurement properties of health status questionnaires.** *J Clin Epidemiol* 2007, **60**:34–42.
15. Fayers PM, Machin D: *Quality of Life: Assessment, Analysis and Interpretation*. Chichester: John Wiley & Sons; 2000.
16. Browne MW, Cudeck R: **Alternative ways of assessing model fit.** *Sociological Methods & Res* 1992, **21**:230–258.
17. Sivo SA, Fan X, Witte EL, Willse JT: **The search for optimal cutoff properties: Fit index criteria in structural equation modeling.** *J Exp Educ* 2006, **74**:267–288.
18. Hu L, Bentler PM: **Cutoff criteria for fit indices in covariance structure analysis: Conventional criteria versus new alternatives.** *Struct Equ Model* 1999, **6**:1–55.
19. Enders CK, Bandalos DL: **The relative performance of full information maximum likelihood estimation for missing data in structural equation models.** *Struct Equ Model* 2001, **8**:430–457.
20. Cheung GW, Rensvold RB: **Evaluating goodness-of-fit indexes for testing measurement invariance.** *Struct Equ Model* 2002, **9**:233–255.
21. Nunnally JC, Bernstein IR: *Psychometric Theory*. 3rd edition. New York: McGraw-Hill; 1994.
22. Benjamini Y, Hochberg Y: **Controlling the false discovery rate: a practical and powerful approach to multiple testing.** *J Royal Statistical Society: Series B* 1995, **57**:289–300. Retrieved <http://www.jstor.org>.
23. Benjamini Y, Krieger AM, Yekutieli D: **Adaptive linear step-up procedures that control the false discovery rate.** *Biometrika* 2006, **93**:491–507.
24. Cohen J: **A Power Primer.** *Psychol Bull* 1992, **112**(1):155–159.
25. Janda M, Steginga S, Langbecker D, Dunn J, Walker D, Eakin E: **Quality of life among patients with a brain tumor and their carers.** *J Psychosom Res* 2007, **63**(6):617–623.
26. Mainio A, Hakko H, Niemelä A, Koivukangas J, Räsänen P: **Gender difference in relation to depression and quality of life among patients with a primary brain tumor.** *Eur Psychiatry* 2006, **21**(3):194–199.
27. Niemelä A, Koivukangas J, Herva R, Hakko H, Räsänen P: **Gender Difference in Quality of Life among Brain Tumor Survivors.** *J Neurol Neurophysiol* 2011, **2**:118.
28. Fischer R, Ferreira MC, Assmar E, Redford P, Harb C, Glazer S, Cheng BS, Jiang DY, Wong C, Kumar N, Kärtner J, Hofer J, Achoui M: **Individualism-collectivism as descriptive norms: Development of a subjective norm approach to culture measurement.** *Journal of Cross-Cultural Psychology* 2009, **40**:187–213.

doi:10.1186/1477-7525-10-54

Cite this article as: Khoshnevisan et al.: Translation and validation of the EORTC brain cancer module (EORTC QLQ-BN20) for use in Iran. *Health and Quality of Life Outcomes* 2012 **10**:54.

Submit your next manuscript to BioMed Central and take full advantage of:

- Convenient online submission
- Thorough peer review
- No space constraints or color figure charges
- Immediate publication on acceptance
- Inclusion in PubMed, CAS, Scopus and Google Scholar
- Research which is freely available for redistribution

Submit your manuscript at
www.biomedcentral.com/submit

